

Lesson 7: Item response theory models (part 2)

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Framework for

e estimating *latent traits* (ability levels) θ

by means of manifest (observable) variables (item responses) and appropriate psychometric (statistical) model

Notes:

- Ability θ is often treated as random variable (but see further)
- **•** Items: dichotomous, polytomous, multiple-choice, ...
- IRT model: describes probability of (correct) answer as function of
	- ability level and
	- item parameters

This function is called:

- \bullet Item response function (IRF)
- Item characteristic curve (ICC)

Use of IRT models

- To calibrate items (i.e. to estimate difficulty, discrimination, guessing,...)
- To assess respondents' latent trait (ability, satisfaction, anxiety,...)
- To describe test properties (standard error, test information,...)
- Test linking and equating, computerized adaptive testing, etc.

IRT model assumptions

¹ Model definition (functional form, usually monotonic ICC)

• e.g. 2PL IRT model:
$$
P(Y_{ij} = 1 | \theta_i, a_j, b_j) = \pi_{ij} = \frac{e^{a_j(\theta_i - b_j)}}{1 + e^{a_j(\theta_i - b_j)}}
$$

- **2** Unidimensionality of latent variable θ
- **3** Local independence (conditional independence)

• e.g.
$$
P(Y_{i1} = 1, Y_{i2} = 1 | \theta_i, a_j, b_j) = \pi_{i1} \cdot \pi_{i2}
$$

• e.g.
$$
P(Y_{i1} = 1, Y_{i2} = 0 | \theta_i, a_j, b_j) = \pi_{i1} \cdot (1 - \pi_{i2})
$$

- **4** Invariance of parameters
- Independence of respondents

$$
\pi_{ij} = P(Y_{ij} = 1 | \theta_i, b_j) = \frac{\exp(\theta_i - b_j)}{1 + \exp(\theta_i - b_j)}
$$
(1)

 θ_i ability of person i, for $i = 1, \ldots, I$ b_i difficulty of item j (location of inflection point) for $j = 1, \ldots, J$

Item Characteristic Curve (ICC)

Note: Originally, Rasch model denoted as $\pi_{ij} = \frac{\tau_i}{\tau_i + \xi_j}$. To get to [\(1\)](#page-4-0), consider $\theta_i = \log(\tau_i)$, and $b_j = \log(\xi_j)$ for $\tau_i > 0, \xi_j > 0$

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Rasch model is sometimes defined as:

$$
\pi_{ij} = P(Y_{ij} = 1 | \theta_i, b_j) = \frac{\exp(D[\theta_i - b_j])}{1 + \exp(D[\theta_i - b_j])}
$$

 $D = 1.702$ is scaling parameter introduced in order to match logistic and probit metrics very closely (Lord and Novick, 1968)

Note: Probit (normal-ogive) model: $\pi_{ij} = \Phi(\theta_i - b_j)$, where $\Phi(x)$ is a cumulative distribution function for the standard normal distribution.

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IRT models allow us to put *items* and *persons* on the same scale

Note: See an example of "32-item test of body height" (van der Linden, 2017), compare to Figure 2.4

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$$
\pi_{ij} = P(Y_{ij} = 1 | \theta_i, a, b_j) = \frac{\exp[a(\theta_i - b_j)]}{1 + \exp[a(\theta_i - b_j)]}
$$

 θ_i ability of person i for $i = 1, \ldots, I$ b_j difficulty of item j (location of inflection point) for $j = 1, \ldots, J$ a discrimination common for all items (slope at inflection point)

$$
\pi_{ij} = P(Y_{ij} = 1 | \theta_i, a_j, b_j) = \frac{\exp[a_j(\theta_i - b_j)]}{1 + \exp[a_j(\theta_i - b_j)]}
$$

 θ_i ability of person i for $i = 1, \ldots, I$ b_j difficulty of item j (location of inflection point) a_j discrimination of item j (slope at inflection point) for $j = 1, \ldots, J$

$$
\pi_{ij} = P(Y_{ij} = 1 | \theta_i, a_j, b_j, c_j) = c_j + (1 - c_j) \frac{\exp[a_j(\theta_i - b_j)]}{1 + \exp[a_j(\theta_i - b_j)]}
$$

 θ_i ability of person i for $i = 1, \ldots, I$ b_i difficulty of item j (location of inflection point) a_i discrimination of item j (slope at inflection point) c_j pseudo-guessing parameter of item j (lower/left asymptote), $j = 1, ..., J$

$$
\pi_{ij} = P(Y_{ij} = 1 | \theta_i, a_j, b_j, c_j, d_j) = c_j + (d_j - c_j) \frac{\exp[a_j(\theta_i - b_j)]}{1 + \exp[a_j(\theta_i - b_j)]}
$$

 θ_i ability of person i for $i = 1, \ldots, I$ b_i difficulty of item j (location of inflection point) a_i discrimination of item j (slope at inflection point) c_i pseudo-guessing parameter of item j (lower/left asymptote) d_i inattention parameter of item j (upper/right asymptote), for $j = 1, \ldots, J$

$$
P(\theta, a_j, b_j, c_j, d_j) = c_j + (d_j - c_j) \frac{\exp[a_j(\theta - b_j)]}{1 + \exp[a_j(\theta - b_j)]},
$$

\n
$$
I_j(\theta, a_j, b_j, c_j, d_j) = \frac{bP}{\delta \theta} = a_j(d_j - c_j) \frac{\exp[a_j(\theta - b_j)]}{\{1 + \exp[a_j(\theta - b_j)]\}^2}
$$

Item information trace lines

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$$
\mathbf{I}(\theta) = \sum_j \mathbf{I}_j(\theta, a_j, b_j, c_j, d_j)
$$

Note: Standard error $SE(\hat{\theta}|\theta) = 1/\sqrt{I(\hat{\theta}|\theta)}$ Reliability $SE(\hat{\theta}|\theta) = \sigma \sqrt{(1 - r_{xx}(\hat{\theta}|\theta))}$

Maximum Likelihood Estimation

Once the data have been collected, we can ask: "Which (item/person) parameters would most likely produce these results?"

Estimating ability parameters:

- Assume five items with known item parameters
- Assume response pattern 11000
- Student with what ability is most likely to produce these responses?

Estimating item parameters:

- Assume 20 students with known abilities $\theta_1, ..., \theta_{20}$
- Assume responses to the first item 11000011110101001110
- Item with what difficulty b is most likely to lead to these student responses?

Problem

- Assume five items (obeying Rasch model) with known item parameters $b_1 = -1.90, b_2 = -0.60, b_3 = -0.25, b_4 = 0.30, b_5 = 0.45.$
- Assume response pattern 11000.
- How likely is average student $(\theta = 0)$ to produce these responses?
- How likely is weaker student ($\theta = -1$) to produce these responses?
- Which student is more likely to produce these responses?

Solution

- **•** calculate probability for each response in the pattern
- 2 calculate probability of the response pattern
	- use assumption of conditional independence: product of probabilities of individual responses in the pattern

Estimating ability parameter θ

Solution

•
$$
P(Y_1 = 1 | \theta = 0) = \frac{e^{(0 - (-1.9))}}{1 + e^{(0 - (-1.9))}}, \dots
$$

•
$$
P(\mathbf{Y} = 11000 | \theta = 0) =
$$

$$
\frac{e^{1.9}}{1+e^{1.9}} \cdot \frac{e^{0.6}}{1+e^{0.6}} \cdot \left(1 - \frac{e^{0.25}}{1+e^{0.25}}\right) \cdot \left(1 - \frac{e^{-0.3}}{1+e^{-0.3}}\right) \cdot \left(1 - \frac{e^{-0.45}}{1+e^{-0.45}}\right) =
$$

 $= 0.87 \cdot 0.65 \cdot 0.44 \cdot 0.57 \cdot 0.61 = 0.086$

•
$$
P(Y = 11000 | \theta = -1) = 0.71 \cdot 0.40 \cdot 0.68 \cdot 0.79 \cdot 0.81 = 0.123
$$

Probabilities of responses 1, 1, 0, 0, 0

Problem

• Student with what ability θ is most likely to produce responses 11000?

Solution

- **O** calculate probability for each response in the pattern (as function of θ)
- 2 calculate probability of the response pattern (as function of θ)
	- **•** this function is known as likelihood function L
	- use assumption of conditional independence:
	- \bullet P(11000| θ) = L(11000| θ) = p₁ · p₂ · (1 p₃) · (1 p₄) · (1 p₅)
- **3** find the maximum value of the likelihood function

Problem

• Student with what ability θ is most likely to produce responses 11000?

Solution

•
$$
P(Y = 11000|\theta) =
$$

\n
$$
\frac{e^{\theta+1.9}}{1+e^{\theta+1.9}} \cdot \frac{e^{\theta+0.6}}{1+e^{\theta+0.6}} \cdot \left(1 - \frac{e^{\theta+0.25}}{1+e^{\theta+0.25}}\right) \cdot \left(1 - \frac{e^{\theta-0.3}}{1+e^{\theta-0.3}}\right) \cdot \left(1 - \frac{e^{\theta-0.45}}{1+e^{\theta-0.45}}\right)
$$

• For which θ is the likelihood the highest?

Log-likelihood

- Reaches maximum for the same θ as likelihood
- **•** Easier to handle

•
$$
\log P(\mathbf{Y} = 11000|\theta) = \log \frac{e^{\theta+1.9}}{1+e^{\theta+1.9}} + \log \frac{e^{\theta+0.6}}{1+e^{\theta+0.6}} + \log \left(1 - \frac{e^{\theta+0.25}}{1+e^{\theta+0.25}}\right) + \log \left(1 - \frac{e^{\theta-0.3}}{1+e^{\theta-0.3}}\right) + \log \left(1 - \frac{e^{\theta-0.45}}{1+e^{\theta-0.45}}\right)
$$

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Maximum Likelihood Estimation - Technical details

Log-Likelihood for response pattern y=11000

For which θ is the likelihood the highest?

- **•** Empirical MLE
	- method of brackets
	- does not provide with standard error of estimate
- Newton-Rhapson
	- looks for $\log L' = 0$ (zero derivative of $\log L$)
	- uses second derivative $\log L''$ to find it quickly: $\theta_{new} = \theta_{old} \frac{\log L'}{\log L''}$
	- derivatives can be further used for estimation of item information and standard error

Problem (Estimating item difficulty b)

 \bullet Assuming that person abilities are known, item with what difficulty b is most likely to produce student responses 110010011000?

Solution

- \bullet calculate probability of student response pattern (as function of b)
	- \bullet this function is again known as likelihood function L
	- use assumption of conditional independence
- **2** find the maximum value of the likelihood function

Note: For 2PL models likelihood-function is 2-dimensional!

Three types of ML Estimates in IRT models

Usually, both **person** and item parameters need to be estimated.

- Joint Maximum Likelihood
	- **•** Used in Winsteps
	- Ping-pong between person and item MLE
	- With increasing number of examinees, number of parameters to be estimated increases
	- May lead to inconsistent, biased estimates
- Marginal Maximum Likelihood
	- Used in IRTPRO, ltm, mirt
	- Assumes *prior* ability distribution (usually $N(0, 1)$)
	- Ability is "integrated out" to get ML estimates of item parameters
	- Expected a posteriori estimates of abilities

Conditional Maximum Likelihood

- **a** Used in eRm
- Only applicable in 1PL (Rasch) models, where:
	- **•** Total score is sufficient statistics for ability
	- Percent correct is sufficient statistics for difficulty

Mathematical and technical details:

$$
L = P(Y|\theta, a, b) = \prod_{i=1}^{I} \prod_{j=1}^{J} \pi_{ij}^{y_{ij}} (1 - \pi_{ij})^{(1 - y_{ij})}
$$

• logarithm simplifies the above expression to sum:

$$
ln L = \sum_{i=1}^{I} \sum_{j=1}^{J} y_{ij} \cdot ln(\pi_{ij}) + (1 - y_{ij}) \cdot ln(1 - \pi_{ij})
$$

- maximization incorporates computation of partial derivatives
- indeterminacy of parameter estimates in the origin and unit
	- person centering
	- item centering

Mathematical and technical details:

• marginal likelihood (θ is integrated out)

$$
L = P(Y) = \int_{-\infty}^{\infty} P(x|\theta, a, b) \cdot g(\theta|a, b) d\theta
$$

- $g(\theta|a, b)$ is so called *prior* distribution (usually assumed N(0,1))
- integration solved using Gauss-Hermite quadrature (numerical integration)

Can be understood as weighted sum: at each theta interval, the likelihood of response pattern rectangle is weighted by that rectangle's probability of being observed

 \bullet L does not depend on θ and can be maximized with respect to a, b

Once the item parameters are known (estimated)

- Maximum likelihood estimator (MLE)
- Weighted likelihood
- Bayes model estimator (BME), maximum a posteriori (MAP)
- Expected a posteriori (EAP)

1 Select set of starting values

- randomly or intelligently
- the closer the starting values are to the actual values the better
- **2** Maximize the likelihood get new estimates
- Check the stopping rule stop if:
	- maximal number of runs is reached
	- likelihood does not change too much

- Log-likelihood the bigger the better
- AIC (Akaike information criterion), BIC (Bayesian information criterion) the smaller the better
- LRT (likelihood ratio test): if significant ($p < 0.05$) submodel is rejected, use model with more parameters
- 3PL model: possibly problems with local maxima, problems to distinguish between models

- Ames $&$ Penfield (2015)
- Comparing ICC of the fitted model to observed proportion of correct responses
- Detection of improbable response patterns
- Comparing number of respondents with given response pattern to what is expected by the model $(X^2 \text{ test})$

Further models

- Polytomous IRT models (ordinal/nominal)
- Multidimensional IRT models
- **Hierarchical IRT models, etc.**
- Accounting for Differential item functioning, etc.

Applications

- Test equating
- Computerized adaptive testing, etc.

- Rasch model, 1PL, 2PL, 3PL, 4PL IRT models
- Item Characteristic Curve (ICC)
- Item Response Function (IRF)
- Item Information Function (IIF)
- Test Information Function (TIF)
- **a** Likelihood function
- Parameter estimation: JML, CML, MML
- Model fit, Item fit, Person fit

Thank you for your attention! <www.cs.cas.cz/martinkova>