Lesson 6: Reliability

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- 1. Introduction
- 2. Estimation Procedures
- 3. Beyond Cronbach's alpha
- 4. More on IRR
- 5. Conclusion

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Classical	test theory			

In behavioral research we are typically interested in the **true score** T but have available only the **observed score** X which is contaminated by some (uncorrelated) **measurement error** e, such that X = T + e.

Examples:

- Admission tests: we are interested in **applicant's knowledge or ability** T, but have available only the test score X
- Grading of essays: We are interested in essay's quality T but we have available only the grader's evaluation X
- Questionnaires on satisfaction: main interest is **respondent satisfaction**, but available are only his/her responses on the questionnaire.

The observed score might vary if we chose different items or different graders.

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Classical	test theory			

Natural questions:

- How much information about the true score is indeed contained in the measurement?
- What is the strength of the relationship between true and observed score?

Reliability theory

- Reliability is defined as squared correlation of the true and observed score $\rho_X=\mathrm{corr}\,^2(T,X)=\rho_{T,X}^2$
- $\rho_X \in \langle 0, 1 \rangle$
- equivalently, reliability can be reexpressed as the ratio of the true score variance to total observed variance $\rho_X = \frac{\operatorname{var}(T)}{\operatorname{var}(X)} = \frac{\sigma_T^2}{\sigma_T^2}$

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Implicatio	ns of low reliabi	ility		

- less accurate estimates of the true score
- wider (less precise) confidence intervals
- need of higher number of subjects to demonstrate differences between groups (keeping the same test power)
- attenuation of correlations, bound of criterion validity

$$\rho_{\scriptscriptstyle X,Y} = \rho_{\scriptscriptstyle T_X, \scriptscriptstyle T_Y} \sqrt{\rho_{\scriptscriptstyle X} \rho_{\scriptscriptstyle Y}} \leq \rho_{\scriptscriptstyle X}$$

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Graphical	interpretation			
				<u>}</u>



Low reliability thus low validity



High reliability and high validity

center of the target represents the value we want to measure

High reliability but low validity

- shots represent independent measurements on one object
- reliability represented by variability of the shots
- validity represented by overall shots' closeness to the center

Observations

- high reliability does not ensure high validity
- validity is bounded by reliability

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Reliability	guidelines			

- Conventional requirement $ho_X \ge .8,$ but see Lee (2012)
 - $\bullet \ \geq .9$ for intelligence tests
 - ullet \geq .7 for personality tests
 - $\sim .6$ for essay marking
- In case of low reliability we should think of instrument revision
 - adding items
 - deleting items
 - in case of graders: training, precise instructions

Importance of proper estimation of reliability

- Overestimation may imply adopting unreliable instrument
- Underestimation may imply (costly) revision of instrument
- Misunderstanding of reliability can imply deletion of important items and lowering validity

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Estimatio	on procedures			

The true score T is not observed, thus we can't estimate reliability from its definition ($\rho_{_{T,X}}^2$ nor σ_T^2/σ_X^2)

Parallel measurements

• equally precise measurements of the same true score:

•
$$X_1 = T + e_1$$
, $X_2 = T + e_2$, $\operatorname{var}(e_1) = \operatorname{var}(e_2) = \sigma_e^2$

ullet the reliability of both measurements is the same ho

• if the errors are uncorrelated, then correlation between the measurements is equal to their (common) reliability $\rho_{X_1,X_2} = \frac{\operatorname{cov}(T+e_1,T+e_2)}{\sqrt{\operatorname{var}(T+e_1)\operatorname{var}(T+e_2)}} = \frac{\sigma_T^2}{\sigma_X^2} = \rho$

The methods differ in how they make use of multiple measurements.

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Estimatio	on procedures			

Use of multiple administrations

Methods employ correlation coefficient btw. observed total scores

- Test-retest method (coefficient of stability)
- Alternate test forms (coefficient of equivalence)

Use of composite measurements

Methods employ correlation coefficient btw. observed partial total scores

- Split-half coefficient
- Average split-half
- Cronbach's aplha (coefficient of internal consistency)

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Test-Ret	est			

- Assumes independent test administrations
 - No memory
 - No improvement between administrations



• Optimal interval 6-12 weeks

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Parallel F	orms			

- Assumes trully paralel forms
 - Equally difficult
 - Parallel items and content
- Assumes the same conditions

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Composit	e measurements			

- Goal is to provide multiple converging pieces of information
- E.g. educational tests, scales, questionnaires, ...

What is the relationship between reliability of composite measurement $X = \sum_{j=1}^{m} X_j$ and reliability of its components?

Spearman-Brown prophecy formula (1910)

Assume X_1, \ldots, X_m parallel measurements (with uncorrelated errors and uncorrelated with true scores). Then reliability of each X_i is the same ρ and the composite reliability is

$$\rho_X = \frac{m \cdot \rho}{1 + (m-1)\rho}$$

Remark: Adding parallel items increases reliability of total score.

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Split-half	coefficient			

- correlation between two subscores corrected for test length
- ullet test is split into two parts, two subscores Y_1,Y_2 are computed

•
$$\rho_{SH} = \frac{2\rho_{Y_1,Y_2}}{1+\rho_{Y_1,Y_2}}$$

- assumes that the two subtests are parallel
- depends on how the split was carried out (even/odd, random,...)
 - even-numbered / odd-numbered
 - with intention to create two halves that are as similar as possible
 - in a random fashion
- we may also compute the mean of all possible split-half coefficients
 - Average split-half

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Cronbach	n's alpha			

• based on idea of splitting the test into individual items

$$\alpha = \frac{m}{m-1} \frac{\sum \sum_{j \neq k} \operatorname{cov} \left(X_j, X_k\right)}{\operatorname{var} \left(X\right)} = \frac{m}{m-1} \left(1 - \frac{\sigma_{X_1}^2 + \dots + \sigma_{X_m}^2}{\sigma_X^2}\right)$$

- popular estimator, provides simple and unique estimation
- equals to composite reliability σ_T^2/σ_X^2 in case of parallel (or at least T-equivalent) items and uncorrelated errors
- in general case and uncorrelated errors, alpha is lower bound to reliability $\alpha \leq \rho_X$ (Novick & Lewis, 1967) and can be viewed as index of internal consistency
- in case of correlated errors, alpha can be lower or greater than reliability

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Cronbach	's alpha - limita	tions		

Cronbach's alpha is a good estimator of reliability for

- parallel (or at least T-equivalent) items and and
- uncorrelated errors

Corrections needed for:

- Correlated errors
 - Example: Reading test, group of items associated with one text.
 - Corrections for correlated errors (Rae, 2006)
- Multidimensional measurement
 - Example: Math test, items measuring arithmetic skills but also reading skills etc.
 - Factor-analysis based estimation of reliability (Raykov & Maurcoulides, 2011)
- More sources of error (multilevel models, G-index)
- Other than normal distribution of item responses (what happens in case of binary items?)

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Beyond C	ronbach's alpha			

How to define and estimate internal consistency in case of binary items?

Martinková P, & Zvára K. Reliability in the Rasch Model. Kybernetika, 43(3), pp. 315-26, 2007. http://www.kybernetika.cz/content/2007/3/315/paper.pdf
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• X_{ij} responses of n students on m items

•
$$X_{ij} = T_i + b_j + e_{ij}$$

• $T_i \sim N(0, \sigma_T^2)$ random, student ability
• b_j fixed, $\sum b_j = 0$, describe item difficulty
• $e_{ij} \sim N(0, \sigma_e^2)$ random error
• total scores $X_i = mT_i + \sum_j b_j + \sum_j e_{ij}$
• reliability: $\rho_X = \frac{var(mT_i)}{var(X_i)} = \frac{m^2 \sigma_T^2}{m^2 \sigma_T^2 + m \sigma_e^2} = \frac{\sigma_T^2}{\sigma_T^2 + \frac{1}{m} \sigma_e^2}$
• Cronbach's alpha:
 $\alpha = \frac{m}{m-1} \frac{\sum_{j \neq k} \text{cov}(X_{ij}, X_{ik})}{var(X_i)} = \frac{m}{m-1} \frac{m(m-1)\sigma_T^2}{m^2 \sigma_T^2 + m \sigma_e^2} = \frac{\sigma_T^2}{\sigma_T^2 + \frac{1}{m} \sigma_e^2}$
• estimate of Cronbach's alpha: $\hat{\alpha} = \frac{m}{m-1} \frac{\sum_{j \neq k} s_{jk}}{\sum_{j,k} s_{jk}}$, where $s_{jk} = \frac{1}{n-1} \sum_{t=1}^n (X_{tj} - \bar{X}_{\bullet j})(X_{tj} - \bar{X}_{\bullet k})$



•
$$SS_T = \sum \sum (\bar{X}_{i\bullet} - \bar{X}_{\bullet\bullet})^2 \sim (m\sigma_T^2 + \sigma_e^2)\chi^2(n-1)$$

• $SS_e = \sum \sum (X_{ij} - \bar{X}_{\bullet j} - \bar{X}_{i\bullet} + \bar{X}_{\bullet\bullet})^2 \sim \sigma_e^2\chi^2((n-1)(m-1))$

Expectations of Mean sums of squares

•
$$\operatorname{E} MS_T = \operatorname{E} SS_T / (n-1) = m\sigma_T^2 + \sigma_e^2$$

•
$$EMS_e = ESS_e/((n-1)(m-1)) = \sigma_e^2$$

Cronbach's alpha

$$\alpha = \frac{\sigma_T^2}{\sigma_T^2 + \frac{1}{m}\sigma_e^2} = \frac{\mathrm{E}\,MS_T - \mathrm{E}\,MS_e}{\mathrm{E}\,MS_T}$$

Cronbach's alpha estimate

$$\hat{\alpha} = \frac{m}{m-1} \frac{\sum \sum_{j \neq k} s_{jk}}{\sum \sum_{j,k} s_{jk}} = \frac{MS_T - MS_e}{MS_T} = 1 - \frac{1}{F}$$



Estimate of Cronbach's alpha can be reexpressed as

$$\hat{\alpha} = \frac{MS_T - MS_E}{MS_T} = 1 - \frac{1}{F}$$

- F statistic used to test the submodel with no subject effect $(H_0:\sigma_T^2=0)$
- Interpretation: alpha close to 1 for F high, i.e. when we reject H_0 , i.e. when admission test well discriminates between students
- Gives confidence intervals
- Estimate is not generally appropriate for more complicated designs

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Logistic a	lpha			

F statistic in

$$\hat{\alpha} = 1 - \frac{1}{F}$$

assumes normality of items

- How does the estimate of reliability behave for binary items?
- Would a new estimate

$$\hat{\alpha}_{log} = 1 - \frac{n-1}{X^2}$$

based on statistic used in similar situation in logistic regression (difference of deviances $X^2 = D(B) - D(A + B)$) give better results for case of binary data?



- Classical model not applicable (binary outcome can't be expressed as sum of T and independent error e)
- IRT models ussually assumed
- Reliability can be defined as (Raykov & Maurcoulides, 2011)

$$\rho_X = \frac{\operatorname{var}\left(\mathrm{E}\left(X|T\right)\right)}{\operatorname{var}\left(\mathrm{E}\left(X|T\right)\right) + \mathrm{E}\left(\operatorname{var}\left(X|T\right)\right)} = \frac{\operatorname{var}\left(\mathrm{E}\left(X|T\right)\right)}{\operatorname{var}\left(X\right)}$$

- Resulting integrals can be evaluated numerically, not explicitly
- Not equal to parallel-forms reliability, but differences negligible (Kim, 2012)
- S-B formula holds only approximately (Martinkova, Zvara 2010)



- Cronbach's alpha is readily applicable also for binary items
- Cronbach's alpha represents generalization of so-called Kuder-Richardson formulae (*Psychometrika*, 1937):

•
$$\hat{\rho}_{KR-20} = \frac{p}{p-1} \left[1 - \frac{\sum \hat{r}_k (1-\hat{r}_k)}{\hat{\sigma}_X} \right]$$
, where \hat{r}_k is easiness of k-th item

• For test with items of common difficulties $\hat{\rho}_{KR-21} = \frac{p}{p-1} \left[1 - \frac{\hat{\mu}(p-\hat{\mu}_k)}{p\hat{\sigma}_X} \right]$, where $\hat{\mu}$ is average total score

Pre-defined values:

- \bullet number of students $n=25,\,50,\,100,\,500$
- number of items $m = 10, \, 20, \, 50, \, 100$
- IRT parameters (difficulty, discrimination, guessing for each item)
- 55 values of σ_T (defines true reliability)
- number of simulates ${\cal N}=1000$

For each combination of n, m and σ_T :

- true reliability computed
- N data sets generated:
 - set of n student abilities generated $T_i \sim \mathrm{N}(0, \sigma_T^2)$
 - Y_{ij} generated from IRT model
 - estimates computed from the data
- $\Rightarrow~N$ estimates $\hat{lpha}_{CR},~{\sf KR}$ -21 and \hat{lpha}_{log}
 - bias and MSE of the estimates plotted out







Bias and MSE of two estimators of reliability, item difficulties from (-0.1, 0.1). Number of students n = 25, number of items m = 10, number of simulates N = 1000.

Bias and MSE of two estimators of reliability, item difficulties from (-3, 3). Number of students n = 25, number of items m = 10, number of simulates N = 1000.

• $\hat{\rho}_{\scriptscriptstyle KR-21}$ is not appropriate in case of different item difficulties







Bias and MSE of two estimators of reliability, number of students $n=25,\,{\rm number}$ of items $m=50,\,{\rm number}$ of simulates N=1000.

Bias and MSE of two estimators of reliability, number of students n=25, number of items m=100, number of simulates N=1000.

• \hat{lpha}_{log} has promising properties especially for high number of items



0.0 0.2 0.4 0.6 0.8 1.0



()

· Cronbach alpha

0.8

1.0

logistic alpha

0.6

n= 100 . m= 10 . N= 1000 0.10 0.05 bias 0.00 teria bant HE -0.05 · Cronbach alpha

> 0.2 0.4 0.6 0.8

0.0 0.2 0.4 0.6 0.8 1.0

-0.10

· logistic alpha

· logistic alpha

reliability

1.0

reliability



reliability

· logistic alpha



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Logistic a	lpha - conclusic	ons		

- Idea behind Logistic alpha was explained
- Logistic alpha has promising properties for some scenarios
- Cases of true reliabilities close to 1 need some adjustment
- Cases of high number of students?

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More on i	nter-rater reliab	oility		

Motivation: Teacher Selection Process



Applicants to classroom job openings in Spokane Public Schools during years (2008/09 - 2012/13)

Motivation: Ratings as Source of Error

54-Pt Screening Rubric:

DATE:	SCREENER:
Job # / Position Title:	
APPLICANT NAME:	
RATING	
	(1-6) enderson to memori data as an ana of attentity
SURLEMING URTERIA 3-4 Sample	nory evidence to support this as an area of strength
	idence to support this cause unus of strangelt
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Washington State Certificate Yes / No	
Required Endersonnel Yes/No	
Rating (1 - 6) 4	
TRAINING	Land for quarticy, algorit and level of constitution additional training militing to the position.
Bating (1 - 6) 4	
EXPERIENCE	May appear to which appearents appear the prediction of statistics with part of counter of parts. A largening conductor could be enter bothy
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Bating (1 - 6) 4	
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Rating (1 - 6) 4	
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	ananded baaring of modules.
Rating (1 - 6) 4	
INTERPERSONAL SKILLS	
	Grudips and musiative effective vectory relationships mith diserve staff, student, parents (parellans, and community
Rating (1 - 6) 4	
CULTURAL COMPETENCY	Look for queple reference to successful strategies for building and maintaining a minimum to wait and maint and
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Conclusion

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Motivatio	n: Questions			

1. Do we select the best applicants?

Do admission ratings predict subsequent teacher quality?

• Goldhaber et al. 2017

2. Can we do better?

What causes error in ratings? How to eliminate the error?

• Martinkova et al. 2015



Mean and range of ratings



Applications ranked by average total score

Are the ratings consistent?



Mean and range of ratings



Applications ranked by average total score

Are the ratings consistent?

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Mean and range of ratings



Applicants ranked by average total score

What is causing the inconsistencies in rating?

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Hiring dat	a: Data structu	re		

- 3986 filled forms
- 1177 applicants
 - internal and external
- 141 raters
 - various levels of experience
- 54 schools
 - 3 school types: elementary, middle, high
- 526 job openings
 - 15 types of jobs: grade teacher, math, English, science, ...

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Model-bas	sed reliability es	timates		

- Estimate IRR while accounting for hierarchical data structure
 - schools, job openings, etc.
 - applicant-school matching, etc.
- Test for possible moderators of IRR
 - internal/external status of the applicant
 - rater experience
- Apply this "model-based IRR" to analyze implications for validity
 - how IRR affects power to predict teacher value added

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Conclusio	bn			

In this presentation, we have

- explained motivation beind reliability
- presented mostly used approaches for reliability estimation
 - test-retest
 - parallel forms
 - split-half coefficient
 - Cronbach's alpha
- presented research on alternative to Cronbach's alpha
- discussed use of model-based reliability estimates (for IRR)

Thank you for your attention! www.cs.cas.cz/martinkova

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