

IMMIGRATION AND INNOVATION: CHINESE GRADUATE STUDENTS IN U.S. UNIVERSITIES

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Immigration and Innovation: Chinese Graduate Students in U.S. Universities*

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Abstract

Immigration is rapidly changing the composition of the R&D workforce in the United States. We study here Chinese chemists and chemical engineers who migrate to the United States for their graduate studies. We analyze productivity at the individual researcher level, thus bypassing the identification issues that earlier studies had to confront when analyzing the relationship between immigration and innovation at the university or firm level. Using new data and measurement techniques, we find robust evidence that Chinese students make disproportionate contributions to the scientific output of their advisors and departments. We attribute this result to a selection effect as it is relatively more difficult for Chinese students to gain admission to U.S. PhD programs. Our results strengthen the case for liberal student migration policies.

JEL Classification: F22, I23, O15, O33, J61

Keywords: High-skilled migration, students, universities, China

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Abstract

Složení pracovní síly ve vede a výzkumu se ve Spojených státech díky imigraci rychle mění. V tomto článku zkoumáme čínské chemiky a chemické inženýry, kteří se stehují do Spojených států za účelem doktorského studia. Analyzujeme produktivitu výzkumných aktivit na individuální úrovni, což nám umožňuje vyhnout se problému s identifikací, který musely řešit předchozí studie zabývající se vztahem mezi imigrací a inovacemi na úrovni univerzit nebo firem. S využitím nových dat a technik měření přinášíme robustní důkaz o disproporcionálním přínosu čínských studentů k vědeckému výstupu jejich vedoucích a kateder. Tento výsledek prisuzujeme efektu selekce, neboť pro čínské studenty je přijetí do amerického doktorského programu relativně obtížnější. Naše výsledky poskytují argument pro liberální přístup k migraci studentů.

1 Introduction

“The first step in winning the future is encouraging American innovation. (...) We’re the home to the world’s best colleges and universities, where more students come to study than any place on Earth. (...) Today, there are hundreds of thousands of students excelling in our schools who are not American citizens.”

(Barack Obama, 2011 State of the Union address)

Immigration is rapidly changing the composition of the R&D workforce in the United States. Immigrants represented 37% of the U.S. Science & Engineering workforce with doctorate degrees in 2000 compared to 24% in 1990 (NSF 2007a). The majority of immigrant scientists and engineers in the U.S. came as students (NSF 2007b). In 2009, 55% of PhD degrees in engineering granted by U.S. universities were to temporary visa holders (NSF 2011), with China the largest source country. In fact, recent PhD graduates from U.S. universities are more likely to have done their undergraduate studies at Tsinghua University or Peking University than at the University of California, Berkeley, or any other institution (Mervis 2008).

The internationalization of the U.S. R&D workforce has important implications for firms and for public policy. On the policy side, the key debate is whether immigration policy for foreign students should be liberal, which ultimately depends on whether the U.S. benefits from the influx of foreign scientists and engineers. Migrants may accelerate the rate of U.S. innovation and strengthen its comparative advantage in graduate-intensive sectors of production (Freeman 2009). However, migration of foreign graduate students may crowd out native students (Borjas 2004) or decrease incentives for natives to engage in scientific careers (Borjas 2006). The latter concerns become more salient if a substantial percentage of highly-skilled migrants return to their home countries after completing their education.

Firms may gain access to engineering and scientific talent not otherwise available by hiring foreign scientists and engineers. However, having a more culturally diverse R&D workforce could potentially have a detrimental effect on performance due to communication or other problems. If the market for scientists and engineers was frictionless, any

such consideration would be fully reflected in their wages. Opportunities for arbitrage may exist, however, in the presence of frictions in the market for scientists and engineers. The only paper we are aware of in the management literature that speaks to these issues is Mithas & Lucas (2010) who show that foreign I.T. workers command a wage premium.

This paper contributes to the growing literature on immigration and innovation. Previous studies have typically analyzed the relationship between the innovative performance of cities (Kerr & Lincoln 2010), states (Hunt & Gauthier-Loiselle 2010), firms (Kerr & Lincoln 2010; Passerman 2011) or universities (Stuen, Mobarak & Maskus 2011; Chellaraj, Maskus & Mattoo 2008; Gurmu, Black & Stephan 2010) and the share of migrants in these entities. The problem with this approach is that the non-random assignment of migrants to cities, states, firms or universities requires the development of sophisticated identification strategies. Here, we directly compare the performance of Chinese students with other students in a given university or lab. Our approach complements the earlier literature and strengthens its findings that migrant scientists and engineers increase the rate of U.S. innovation.

To enable an analysis at the level of the individual student, we collected and merged data from multiple unconventional sources. Dissertation abstracts were used to build lists of students.¹ We identified Chinese students through an ethnic name matching algorithm as in Kerr (2008a; 2008b). Information on advisors was retrieved from the faculty listings of the American Chemical Society Directory of Graduate Research. Finally, we matched both the students and their advisors to their publications. Our final dataset covers more than 20,000 PhD students graduating from U.S. universities in chemistry, biochemistry and chemical engineering. These fields account for almost a fourth of PhD degrees in science and engineering granted in the U.S. in 2006 (NSF 2009).

Regressing different measures of productivity on an indicator variable for Chinese student, we find that Chinese students significantly outperform other students. The magnitude of the differential (between +24% and +62%) is sensitive to the choice of productivity measure but robust to the inclusion of either school fixed effects or advisor fixed effects. Next, we compare Chinese students to three other interesting groups of students: Indian students, Korean students, and awardees of the NSF doctoral fellowship program-

¹The use of dissertation abstracts was inspired by the work of Kahn & MacGarvie (2011; forthcoming).

America’s best and brightest in Science and Engineering. Conditional on acceptance into the same program or working with the same advisor, Chinese students perform about as well as the NSF fellows. We also find that the Chinese students outperform both the Indian students and the Korean students. Thus, our results point to the exceptionalism of China with respect to other source countries.

We also run a set of regressions where the unit of analysis is the advisor. We regress the productivity of the advisor as dependent variable on the number of Chinese students and the number of other students graduating that year, the subsequent year and preceding year. Chinese students seem to have a larger effect on the productivity of their advisor than non-Chinese students, which is consistent with the results obtained at the individual student level.

Our preferred explanation for the productivity differential between the Chinese students and other students is a selection effect. Evidence from Attieh and Attieh (1997) suggests that top U.S. universities give substantial preference to U.S. citizens in their admissions decisions. This could reflect a preference for natives but it may also be the consequence of greater uncertainty regarding applicants from other countries. In either case, the bar to admissions is effectively higher for foreign students. The fact that Chinese migrants overwhelmingly come from a small number of elite Chinese universities provides strong support for the importance of these selection processes.

The rest of the paper is organized as follows. We first provide background information on the migration of Chinese students, on the National Science Graduate Fellowships, and on graduate study and knowledge production in chemistry. We then describe our data and present the results. Finally, we discuss potential explanations for the productivity differential between Chinese and other students and conclude.

2 Background

2.1 Migration of Chinese students

The U.S. Immigration and Nationality Act of 1952 formalized the status of foreign students in U.S. institutions by creating “non-immigrant” visa categories permitting tem-

porary residence for study purposes (Bound et al. 2009). The “F1” visa grants access to the US to students who are admitted by a recognized academic institution and who can prove sufficient financial support. The number of student visas issued by the State Department climbed sharply from 65,000 in 1971 to 315,000 in 2000 (U.S. Department of State, Annual Report of the Visa Office, various years).

Until 1978, Chinese migration policy only allowed migration into socialist countries. The relaxation of this policy began a period of steadily rising Chinese migration to the US. In 2002, there were more than 700,000 temporary immigrants to the U.S. from China, about one-tenth of whom were students (USCIS data reported in Poston & Luo (2007)). Around 325,000 scientists and engineers based in the US in 2003 were born in China, Hong-Kong or Macau (NSF 2007b), of whom around three-fourths had obtained at least one university degree in the US (ibid.).

Since the end of the Cultural Revolution in 1977, China has made considerable investments in its higher education system, both in absolute terms and relative to other countries at similar levels of development.² From 1978 to 2006, the number of institutions of higher education more than tripled (Li, 2009) and enrollments increased even faster, growing at approximately 30% a year since 1999 (Li et al. 2008). In the 90’s, Chinese universities graduated slightly fewer Bachelors in Science and Engineering than U.S. universities but the number of Bachelors in S/E graduating from Chinese universities rose sharply in the 00s.

Until recently, Chinese universities offered limited possibilities for graduate education.³ Conversely, the leadership position of US universities as providers of high-quality doctoral education is undisputed. Chinese scientists and engineers report educational opportunities as the most important reason for coming to the United States (NSF 2007b).

²Resource inputs have also been concentrated on a small number of ‘elite’ institutions and in Science and Engineering departments in particular.

³Only slightly more than 1000 doctoral degrees in S/E were awarded in China in 1989 (NSF 2007a). The number of doctoral degrees in S/E awarded by Chinese universities reached 12,000 in 2003 (ibid), which was still only about one third of the the number of doctoral degrees in S/E awarded by US universities in the same year.

2.2 The NSF graduate research fellowship program

Although we are not interested in the National Science Foundation (NSF) graduate research fellowship (GRF) program per se, we provide a brief description as the NSF graduate fellows are a useful reference group for assessing the performance of Chinese students. The NSF graduate research fellowship is a highly prestigious award for Science and Engineering students. Freeman et al. (2005) refer to the program as 'Supporting the "Best and Brightest" in Science and Engineering'. Applicants to GRF program have to be either US citizens or permanent residents to be eligible. Around 1,000 fellowships are awarded each year, which amounts to two fellowships per thousand Science and Engineering Bachelors. Applications are evaluated by panels based upon recommendation letters, graduate point average (GPA) obtained in undergraduate studies and quantitative and verbal graduate research examination (GRE) scores. The program provides financial support for three years of graduate study. While the dollar value of the stipend was relatively low in the 90's (at USD 15,000), the prestige of the award is considerable and could easily be leveraged to obtain the best possible financial support from host institutions, as universities actively woo NSF graduate fellows (WestEd 2002).

2.3 Graduate study and knowledge production in chemistry

As in much of the physical and life sciences, knowledge production in chemistry is organized in laboratories. Typically, graduate students focus on conducting experiments while the faculty member focuses on conceptual-theoretical activities, including raising grant funding, formulating research questions, designing research projects and interpreting experimental results (Laudel 2001). In the words of the Economist (2007) the graduate student is 'the workhorse of the modern laboratory'. Correspondingly, faculty and students work in much closer collaboration than in the social sciences. Authorship practices reflect the division of labor within the lab. The median number of authors per paper is five with the first authorship conferred on the junior scholar who did most of the experiments. The name of the advisor normally appears at the end of the authorship list and more minor contributions to the paper are recognized through authorship slots in the middle. Graduate students are not expected to publish independently of their faculty advisors.

The median enrollment to graduation time is six years for chemistry PhD students (NSF 2006). In the first year of PhD graduate programs, students choose an advisor in whose lab they will conduct research, in principle for the rest of their PhD study. They are mainly supported by research assistantships (42.2%) and teaching assistantships (41.8%) while around 8% have fellowships (NSF 2007c). About half of graduating PhD students pursue careers in industry, either in the classical chemical industry or, increasingly, in the pharmaceutical industry. About 30% of chemistry PhD graduates pursue careers in the academia.

3 Data

We identify PhD students using Proquest Dissertations and Abstracts. This bibliographic database lists abstracts of completed PhD theses with the name of the student, the university and year of graduation as well as the field and the name of the advisor. It also includes links to the full-text of the theses, which is useful because theses from certain universities include additional bibliographical information on students (MacGarvie 2007).

Proquest data provides good coverage of PhD graduates for U.S. universities and recent years. However, we do not directly observe country of birth or of undergraduate education. To address this limitation, we use a technique similar to that pioneered by Kerr (2008a, 2008b) in his study of ethnic patent inventors. This technique relies on the fact that names implicitly contain information about the origin of individuals and ethnicity can be reasonably well inferred by matching names to lists of ethnic names. Using the same approach, we constructed lists of Chinese last and first names and used them to code students as Chinese.

To verify the quality of the results obtained with of our ethnic name matching algorithm, we manually coded hundreds of CVs for universities that require students to report biographic information in theses. We find that 88% of students coded as Chinese had received their undergraduate degrees in China (and a further 5% in Taiwan). Conversely, our algorithm identified 96% of students who did their undergraduate studies in China as Chinese (see table 1).

[insert table 1 about here]

We construct scientific output measures by matching our list of students to publication data from Scopus, a bibliographic database from Elsevier. The matching is by no means trivial and its validity is essential to the credibility of our exercise. To minimize errors in the matching process, we exploit the fact that papers authored by chemistry students are almost invariably written with their advisor as coauthor, as discussed earlier. We also use affiliation data for individual authors.⁴ A publication is matched to a student if nine criteria are successfully met: an author of the publication needs to have the (1) last name of the student (2) first initial of the student (3) correct departmental affiliation of the student (4) correct university affiliation of the student; and one of the coauthors of the paper has to have (5) the last name of the advisor (6) the first initial of the advisor (7) the correct departmental affiliation of the advisor and (8) the correct university affiliation of the advisor. Finally, the paper must have been published (9) no earlier than 3 years prior to the graduation of the student and no later than the year of graduation.

Scientific output can be measured by counting the number of publications, or adjusting for the journal impact factor (JIF) (a reasonable proxy for quality of the journal) or by counting the number of cites received by these publications. Moreover, as discussed earlier, first-authorship in chemistry has a special meaning, being typically used to recognize the junior scholar who made the main contribution to the paper. Thus, it is interesting to distinguish between first-authored publications and all publications. Although our preferred measure of productivity is counts of first-authored articles weighted by journal impact factor, we use the other productivity measures as well.

Proquest includes names of PhD advisors but no other information on them. To enrich our data set, we match advisors from Proquest to the faculty listings of the Directory of Graduate Research from the American Chemical Society. This data source has information on the age, gender as well as educational and professional histories of chemistry and chemical engineering faculty.⁵ Finally, we infer the area of specialization of faculty in our sample from the journals in which they publishes. For instance, a faculty member who often publishes in the *Journal of Biological Chemistry* is assumed to specialize in

⁴Scopus includes affiliation data for individual authors which is not the case for the concurrent bibliographic database ISI Web of Science.

⁵For more information about this database, see Gaule (2011)

biochemistry. Table 2 displays a listing of areas of specialization and an example journal for each.

[insert table 2 about here]

4 Descriptive statistics

Our data cover virtually all U.S. PhD-granting departments in chemistry and chemical engineering. We have a total of 21,154 students graduating between 1999 and 2008 of whom 2,220 (10.49%) are identified as Chinese by our name matching algorithm. The share of Chinese students is slightly increasing over time from an average of 9.5% for graduation years from 1999 to 2003 to an average of 11.6% for 2004 to 2008. The share of Chinese students exhibits considerable variation across universities. For instance, more than 50% of graduates from New York University are Chinese compared to less than 3% at the University of California, Berkeley. The percentage of Chinese students is markedly higher in schools with lower chemistry R&D budgets. For instance, Chinese students represent 7.2% of students in the top 25 chemistry departments compared to 12.6% for the rest of the departments.

5,139 faculty advised at least one student graduating between 1999 and 2008, with a mean of 4 students and a maximum of 29. An interesting pattern that emerges in our data is that Chinese students are more likely to match with Chinese advisors.⁶ 29.0% of students advised by Chinese faculty in U.S. universities are themselves Chinese though the Chinese students represent only 10.5% of the student population. However, we have only 115 Chinese advisors in our sample (less than 3% of advisors) so that only 5.2% of Chinese students graduate with a Chinese advisor.

Finally, descriptive statistics by type of student for the six output measures are displayed in table 3. From this raw data, we can see that Chinese students have higher unconditional mean productivity than other students for the six output measures.

[insert table 3 about here]

⁶Similar patterns for a smaller sample are found by Tanyildiz (2008)

5 Estimation and results

5.1 Regressions at the individual student level

In the first step of our analysis, we regress various measures of scientific output on an indicator variable for Chinese students. Our controls always include sets of indicator variables for the year of graduation and for the specialization of the advisor. We alternate between specifications with school fixed effects, with advisors fixed effects and with neither.

Most regressions are estimated with a quasi-maximum likelihood conditional fixed-effects Poisson model (“Poisson QML”; see Hausman et al. 1984). This model has several desirable properties, including consistency of the coefficient estimates independently of any assumption on the conditional variance as long as the mean is correctly specified (Woolridge 1997) and consistency in the standard errors even if the data generating process is not Poisson. This estimator can also be used for fractional and non-negative variables (Santos Silva & Tenreyro 2006), such as publications counts adjusted by journal impact factors in our case. We implement this in Stata with the “xtqmlp” procedure written by Tim Simcoe.⁷ Poisson QML estimates are interpreted as $(\exp(\beta) - 1) * 100$ percentage change or approximately as $\beta * 100$ percentage change.⁸ They can thus be easily compared across regressions.

The results are displayed in table 4 . Each panel (A, B, C) presents the results obtained with different sets of fixed effects. The coefficient for Chinese student is positive and significant at 1% in all specifications. Comparing the coefficient for Chinese student horizontally across different productivity measures, the coefficient tends to be larger for citations than for raw publication counts, or counts weighted by journal impact factors. Since citations are usually thought to be a better proxy for the quality of research, this suggests that the productivity advantage of Chinese students is stronger in the quality dimension than in the quantity dimension. Comparing the Chinese student coefficient vertically across panels, we find that it is not sensitive to the inclusion of either school fixed effects or advisor fixed effects. Thus, the productivity differential of Chinese students

⁷Available for download at <http://people.bu.edu/tsimcoe/code/xtpqml.txt>

⁸For instance, a coefficient of 0.216 corresponds to $\exp(0.216) - 1 = 0.241 * 100 = 24.1\%$ with $0.216 * 100 = 21.6\%$ a decent approximation.

does not seem to be explained by school or advisor characteristics. The point estimates range from 0.216 to 0.497, implying that Chinese students are between 24.1% ($100 * (\exp(0.216) - 1)$) and 64% ($100 * \exp((0.497) - 1)$) more productive than other students.

[insert table 4 and 5 about here]

So far, we have compared Chinese students to all other students combined. However, it is also interesting to compare Chinese students to other, more specific, populations of students. We thus repeat the same type of exercise and introduce indicator variables for Indian students, Korean students⁹, and Fellows of the National Science foundation (an elite group of American students). The results are displayed in panels D, E, and F of table 5. The results for the other groups of students are somewhat less stable across specifications than those for Chinese students, perhaps due to the smaller size of these groups. Nevertheless, clear patterns emerge. While Indian and Korean students outperform the residual group of students who are neither Chinese, Indian, Korean or NSF fellows, they systematically underperform the Chinese students. Another interesting comparison is between the coefficient for Chinese students and the NSF fellows. Conditional on being in the same school (panel E) or having the same advisor (panel F), the Chinese students perform about as well the NSF fellows- America's best and brightest in Science and Engineering.

5.2 Regressions at the individual advisor level

We continue our analysis by using a completely different, but complementary, approach. We use the information we have on students and advisors to construct a panel of advisors over the years 2000-2007. We then regress the output of the advisor on the number of Chinese and other students graduating in that year, the preceding year, and subsequent year. The estimation is by ordinary least squares with robust standard errors. Our controls include school fixed effects, advisor age fixed effects and specialization fixed effects. As discussed earlier, a chemistry faculty member typically relies on her graduate

⁹Indian and Korean students are identified using an ethnic name matching algorithm, like the Chinese students.

students to conduct experiments that she conceives, and her name would invariably be included in the resulting publications.

[insert table 6 about here]

The results are shown in table 6. Having an extra Chinese student in the lab increases the output of the advisor by 0.895 paper per year on average whereas a non-Chinese student increases the output of the advisor by 0.630 only. A F-test rejects the equality between the coefficient for Chinese students coefficient and the coefficient for other students at the 1% confidence level. Similar results are obtained when using citations or publications weighted by journal impact factor. Thus, Chinese students seem to make a disproportionate contribution to the output of their advisors.

The results of these regressions should be interpreted with caution because graduate students are typically financed through grants to their advisors. Thus, the number of students that a faculty member has will reflect his (unobserved) past fund-raising success. Furthermore, the assignment of students to advisors may be non-random. While these considerations may threaten the consistency of point estimates for either type of students, it is less clear that they affect the difference between the two types of students. Thus, while the results at the advisor level are entirely consistent with those at the individual student level, we place more confidence in the individual student regressions.

6 Discussion

Why do the Chinese students perform well? Our preferred explanation is a selection effect. U.S. education enjoys an excellent reputation in China and attracts the brightest and most motivated Chinese students. Despite the fact that U.S. universities are admitting large numbers of Chinese students, it is nevertheless considerably more difficult for a Chinese to be accepted into a U.S. PhD program. Evidence from Attieh and Attieh (1997) suggests that top U.S. universities give substantial preference to U.S. citizens in their admission decisions. While this may reflect an underlying preference for natives, it could also be an optimal response to difficulties encountered in evaluating the applications of Chinese students (lack of familiarity with schools, grading systems and reference letter writers).

Precisely because of these difficulties, an undergraduate degree from one of the top Chinese universities is a *de facto* requirement for entry into a U.S. PhD program. Chinese graduate students come from a restricted set of elite Chinese universities. In China, around 10 million high school finishers take the national college entrance exam but only three thousand (0.03%) are admitted into the two most prestigious schools, Peking University and Tsinghua University. Peking University and Tsinghua University are thus more selective than the most exclusive U.S. institutions; the majority of MIT undergraduates would not have had standardized test scores high enough to be admitted into the undergraduate programs of Peking University and Tsinghua University.¹⁰

Besides this positive selection story, two other plausible arguments could explain the productivity advantage of the Chinese students. The first argument points to incentive effects due to higher preferences of the Chinese students for an academic career. The second explanation emphasizes a cultural predisposition of the Chinese students for higher effort at work.

According to the first argument, a career in academia and a post-doctoral training in particular may be relatively more attractive to Chinese students, thus increasing the incentives to publish during PhD study. In particular, immigration considerations may be relevant as maintaining valid immigration status in the U.S. could be easier when undertaking post-doctoral training than when working in industry.¹¹ However, data from the 2001 Survey of Earned Doctorates indicate that Chinese PhD students in chemistry are more likely than others to be planning to go to industry directly from graduate school (35.8% of those born in China plan to work in industry, versus only 26.2% of those born anywhere else) and no more likely to be planning to follow postdoctoral training (37.4% of the Chinese plan to follow postdoctoral training, versus 39.7% of all other students) (Stephan 2010).

In support of the second argument, anecdotal evidence suggests that Chinese graduate students work harder and spend more time in the laboratory. Again, complementary evidence from other data sources suggests that this cultural predisposition for higher

¹⁰The median SAT math score of MIT undergraduates is 770 which is lower than the top centile cutoff of the SAT score distribution.

¹¹From the perspective of immigration law, post-doctoral training is not considered work. Most post-doctoral fellows are on visitor (J1) rather than on work (H1B) visas. The latter, but not the former, are subject to a yearly cap.

effort might be of second order importance. In fact, a survey of post-docs found only small differences between Chinese and Americans in terms of hours worked (50.5 hours per week versus 49.8; Brumfiel 2005).

7 Concluding remarks

The contribution of this paper is to show that Chinese students enrolled in U.S. PhD programs have a strong record of publications during their graduate studies and make disproportionate contributions to the productivity of their advisors and departments. We argue that this excellent performance of the Chinese in comparison to other students is most likely explained by selection processes. Our findings strengthen and complement other studies that have found that migrant scientists and engineers make disproportionate contributions to U.S. innovation (Levin & Stephan 1999; Hunt 2011, Stuen, Mobarak & Maskus 2011; Chellaraj, Maskus & Mattoo 2008, Gurmu, Black & Stephan 2010). The influx of talented students is an important benefit to host universities. However, potential negative effects, and in particular the fact that migration may decrease incentives for natives to engage in scientific careers (Borjas 2006), should also be taken into account.

Since the costs of training graduate students are mainly borne by the host country in the form of research and teaching assistantships, it is important to consider the stay rates of Chinese students. According to estimates derived by Finn (2007) using Social Security data, the stay rate for Chinese doctorate recipients is around 92% after five years from obtaining the title, the highest observed for any major source country in 2005. Similarly, Gaule (2014) finds that the lifetime odds of permanent return for Chinese migrants who become faculty in U.S. universities is less than 2%. However, the current high stay rates cannot be taken for granted given the steadily rising skill premium in China and the aggressive recruiting policies of Chinese universities.

An important limitation of our study is that our research is based on an early measure of productivity - publication during the PhD - and does not address post-PhD outcomes. We do not know if Chinese students continue to outperform non-Chinese students after their graduation. On the one hand, initial differences in scientific productivity tend to persist and might even amplify over time. On the other hand, post-graduate occu-

pations may require skills that the Chinese immigrants may lack. Further evidence on productivity and mobility choices after the PhD would thus be useful to complete our understanding of the contribution of high-skilled Chinese migrants to U.S. innovation.

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Figures

Figure 1: Illustrating the matching process

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by [Huang, Yu](#), Ph.D., [Harvard University](#), 2003, 139 pages; AAT 3106649

Science 9 November 2001:
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REPORTS

Logic Gates and Computation from Assembled Nanowire Building Blocks
[Yu Huang](#),^{1*} [Xiangfeng Duan](#),^{1*} [Yi Cui](#),¹ [Lincoln J. Lauhon](#),¹ [Kyoung-Ha Kim](#),² [Charles M. Lieber](#)^{12†}

Miniaturization in electronics through improvements in established "top-down" fabrication techniques is approaching the point where fundamental issues are expected to limit the dramatic increases in computing seen over the past several decades. Here we report a "bottom-up" approach in which functional device elements and element arrays have been assembled from solution through the use of electronically well-defined semiconductor nanowire building blocks. We show that crossed nanowire p-n junctions and junction arrays can be assembled in over 95% yield with controllable electrical characteristics, and in addition, that these junctions can be used to create integrated nanoscale field-effect transistor arrays with nanowires as both the conducting channel and gate electrode. Nanowire junction arrays have been configured as key OR, AND, and NOR logic-gate structures with substantial gain and have been used to implement basic computation.

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Proquest Dissertation and Abstracts (upper part of the figure) has information on the name of the student, name of the advisor, university and year of graduation. The publication (lower part of the figure) has information on the affiliation of each author. A publication is matched to a student if nine criteria are successfully met: The first author of the publication needs to have (1) last name of the student (2) first initial of the student (3) correct departmental affiliation of the student (4) correct university affiliation of the student; and one of the coauthors of the paper had to have (5) the last name of the advisor (6) the first initial of the advisor (7) the correct departmental affiliation of the advisor and (8) the correct university affiliation of the advisor. Finally, the paper had to be published (9) no earlier than 3 years prior to the graduation of the student and no later than the year of graduation.

Tables

Table 1: **Using biographic information to verify the quality of the name matching**

With Chinese name	
<i>Right Match</i>	88.2% are educated in China
<i>Wrong Match</i>	11.8% not educated in China (5.1% Taiwan)
Educated in China	
<i>Right Match</i>	95.6% have a Chinese name
<i>Wrong Match</i>	4.4% do not have a Chinese name

Table 2: **Using journals to define field of specialization**

Field	Journal (example)
Analytical chemistry	<i>Analytical Chemistry</i>
Applied Chemistry	<i>Angewandte Chemie</i>
Biochemistry	<i>Journal of Biological Chemistry</i>
Chemical Engineering	<i>AIChE Journal</i>
Inorganic Chemistry	<i>Inorganic Chemistry</i>
Material Science	<i>Macromolecules</i>
Organic Chemistry	<i>Journal of Organic Chemistry</i>
Nanotechnology	<i>Nano Letters</i>

The area of specialization for a given faculty member is inferred from the journals s/he publishes in. For instance, a faculty member who publishes often in the *Journal of Biological Chemistry* is assumed to be specialized in biochemistry.

Table 3: **Descriptive statistics on scientific output**

	Chinese students (n=2,220)		All other students (n=18,934)	
	Mean	SD	Mean	SD
First-authored publications				
Number	1.16	1.47	0.93	1.27
JIF-weighted	5.57	8.25	4.63	7.38
Cites to	34.74	131.19	27.19	83.48
All publications				
Number	1.81	2.48	1.41	2.03
JIF-weighted	9.29	15.37	7.25	12.24
Cites to	66.63	264.45	46.01	139.88

Table 4: **Regression results at the individual student level**

(A)	First-authored publications			All publications		
	Number	JIF-weighted	Cites to	Number	JIF-weighted	Cites to
Chinese student	0.216*** (0.021)	0.189*** (0.010)	0.275*** (0.004)	0.247*** (0.017)	0.250*** (0.007)	0.396*** (0.003)
Year of graduation FE	Yes	Yes	Yes	Yes	Yes	Yes
Lab specialization FE	Yes	Yes	Yes	Yes	Yes	Yes
Students	21,154	21,154	21,154	21,154	21,154	21,154

(B)	First-authored publications			All publications		
	Number	JIF-weighted	Cites to	Number	JIF-weighted	Cites to
Chinese student	0.237*** (0.033)	0.248*** (0.038)	0.370*** (0.115)	0.265*** (0.034)	0.311*** (0.042)	0.497*** (0.122)
Year of graduation FE	Yes	Yes	Yes	Yes	Yes	Yes
School FE	Yes	Yes	Yes	Yes	Yes	Yes
Lab specialization FE	Yes	Yes	Yes	Yes	Yes	Yes
Students	21,133	21,133	21,133	21,133	21,133	21,133
Universities	164	164	164	164	164	164

(C)	First-authored publications			All publications		
	Number	JIF-weighted	Cites to	Number	JIF-weighted	Cites to
Chinese student	0.231*** (0.027)	0.239*** (0.032)	0.297*** (0.085)	0.236*** (0.029)	0.269*** (0.034)	0.390*** (0.090)
Year of graduation FE	Yes	Yes	Yes	Yes	Yes	Yes
Advisor FE	Yes	Yes	Yes	Yes	Yes	Yes
Students	17,185	17,183	17,115	17,185	17,183	17,141
Advisors	3,013	3,012	2,991	3,013	3,012	3,000

Robust standard errors in parentheses. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

Note: Panel A specifications are estimated using a Poisson regression with robust standard errors. Panel B and C specifications are estimated by Poisson Quasi-Maximum Likelihood. The various productivity measures used as dependent variable are based on papers published between three years before graduation and the year of graduation.

Table 5: **Regression results at the individual student level - continued**

(D)	First-authored publications			All publications		
	Number	JIF-weighted	Cites to	Number	JIF-weighted	Cites to
Chinese student	0.234*** (0.021)	0.212*** (0.010)	0.300*** (0.004)	0.264*** (0.017)	0.270*** (0.008)	0.414*** (0.003)
Indian student	0.104*** (0.030)	0.079*** (0.014)	0.027*** (0.006)	0.065** (0.025)	0.028** (0.011)	-0.065*** (0.005)
Korean student	0.065* (0.035)	0.108*** (0.015)	0.112*** (0.006)	0.049* (0.029)	0.075*** (0.012)	0.074*** (0.005)
NSF fellow	0.461*** (0.045)	0.654*** (0.018)	0.846*** (0.007)	0.499*** (0.036)	0.653*** (0.014)	0.803*** (0.005)
Year of graduation FE	Yes	Yes	Yes	Yes	Yes	Yes
Lab specialization FE	Yes	Yes	Yes	Yes	Yes	Yes

(E)	First-authored publications			All publications		
	Number	JIF-weighted	Cites to	Number	JIF-weighted	Cites to
Chinese student	0.256*** (0.034)	0.269*** (0.039)	0.385*** (0.108)	0.283*** (0.035)	0.328*** (0.043)	0.502*** (0.114)
Indian student	0.130*** (0.042)	0.137*** (0.045)	0.104 (0.066)	0.094** (0.043)	0.090* (0.047)	0.022 (0.065)
Korean student	0.078* (0.047)	0.111* (0.060)	0.106 (0.120)	0.055 (0.044)	0.068 (0.057)	0.047 (0.098)
NSF fellow	0.337*** (0.043)	0.301*** (0.043)	0.162 (0.170)	0.347*** (0.047)	0.256*** (0.045)	0.087 (0.110)
Year of graduation FE	Yes	Yes	Yes	Yes	Yes	Yes
School FE	Yes	Yes	Yes	Yes	Yes	Yes
Lab specialization FE	Yes	Yes	Yes	Yes	Yes	Yes

(F)	First-authored publications			All publications		
	Number	JIF-weighted	Cites to	Number	JIF-weighted	Cites to
Chinese student	0.245*** (0.027)	0.256*** (0.032)	0.316*** (0.085)	0.246*** (0.029)	0.278*** (0.034)	0.401*** (0.092)
Indian student	0.145*** (0.039)	0.151*** (0.046)	0.160** (0.072)	0.094** (0.039)	0.076 (0.046)	0.054 (0.063)
Korean student	0.040 (0.043)	0.055 (0.059)	0.088 (0.124)	-0.018 (0.045)	-0.030 (0.060)	-0.024 (0.094)
NSF fellow	0.318*** (0.053)	0.335*** (0.060)	0.265** (0.125)	0.359*** (0.055)	0.323*** (0.060)	0.252** (0.114)
Year of graduation FE	Yes	Yes	Yes	Yes	Yes	Yes
Advisor FE	Yes	Yes	Yes	Yes	Yes	Yes

Robust standard errors in parentheses. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

Note: Panel E specifications are estimated using a Poisson regression with robust standard errors. Panel E and F specifications are estimated by Poisson Quasi-Maximum Likelihood. The various productivity measures used as dependent variable are based on papers published between three years before graduation and the year of graduation.

Table 6: **Regressions at the individual advisor level**

	Publications of the advisor		
	Number	JIF-Weighted	Cites to
Number of Chinese students	0.895*** (0.086)	4.641*** (0.565)	30.557*** (7.497)
Number of other students	0.630*** (0.031)	3.357*** (0.214)	18.604*** (1.854)
School fixed effects	Yes	Yes	Yes
Speciality fixed effects	Yes	Yes	Yes
Age fixed effects	Yes	Yes	Yes
Observations	34,005	34,005	34,005
Advisors	4,870	4,870	4,870

Robust standard errors in parentheses. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

We regress the output of the advisor on the number of Chinese and other students graduating in that year, the preceding year, and subsequent year. The panel is at the advisor-year level and runs from 2000 to 2007.

Table 7: **Ethnic labs**

Chinese students	as % of all students 10.5%	as % of students with Chinese advisor 29.0%
Indian students	as % of all students 5.8%	as % of students with Indian advisor 17.3%
Korean students	as % of all students 4.1%	as % of students with Korean advisor 32.3%

Table 8: **Ethnic labs - continued**

Students with Chinese advisor	as % of all students 1.9%	as % of Chinese students 5.2%
Students with Indian advisor	as % of all students 2.4%	as % of Indian students 7.2%
Students with Korean advisor	as % of all students 0.5%	as % of Korean students 3.6%

Table 9: Interaction between Chinese student and Chinese advisor

(G)	First-authored publications			All publications		
	Number	JIF-weighted	Cites to	Number	JIF-weighted	Cites to
Chinese student	0.219*** (0.034)	0.229*** (0.038)	0.338*** (0.126)	0.238*** (0.036)	0.284*** (0.044)	0.482*** (0.144)
Chinese advisor	0.561*** (0.093)	0.679*** (0.107)	1.134*** (0.203)	0.781*** (0.112)	0.904*** (0.145)	1.541*** (0.279)
Chinese student with Chinese advisor	0.031 (0.107)	-0.024 (0.139)	0.051 (0.301)	0.010 (0.102)	-0.030 (0.151)	-0.262 (0.308)
Year of graduation FE	Yes	Yes	Yes	Yes	Yes	Yes
School FE	Yes	Yes	Yes	Yes	Yes	Yes
Lab specialization FE	Yes	Yes	Yes	Yes	Yes	Yes
Students	21,133	21,133	21,133	21,133	21,133	21,133
Universities	164	164	164	164	164	164

(H)	First-authored publications			All publications		
	Number	JIF-weighted	Cites to	Number	JIF-weighted	Cites to
Chinese student	0.231*** (0.028)	0.236*** (0.033)	0.270*** (0.085)	0.242*** (0.029)	0.271*** (0.034)	0.397*** (0.094)
Chinese student with Chinese advisor	0.016 (0.127)	0.040 (0.156)	0.288 (0.291)	-0.046 (0.142)	-0.022 (0.150)	-0.054 (0.202)
Year of graduation FE	Yes	Yes	Yes	Yes	Yes	Yes
Advisor FE	Yes	Yes	Yes	Yes	Yes	Yes
Students	17,185	17,183	17,115	17,185	17,183	17,141
Advisors	3,013	3,012	2,991	3,013	3,012	3,000

Robust standard errors in parentheses. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

Note: Panel G and H specifications are estimated by Poisson Quasi-Maximum Likelihood. The various productivity measures used as dependent variable are based on papers published between three years before graduation and the year of graduation.

Table 10: **Baccalaureate-origin institutions of PhD graduates in S/E from US Universities**

Baccalaureate-origin Institution	Country	Rank	All S/E doctorates	Life sciences	Physical Sciences	Eng.
Tsinghua Univ	China	1	542	17	104	421
Beijing Univ	China	2	435	139	221	75
Seoul National Univ	Korea	3	239	56	76	107
Cornell Univ	USA	4	210	108	58	44
Univ of California-Berkeley	USA	5	207	92	59	56
National Taiwan Univ	Taiwan	6	176	64	49	63
Massachusetts Inst of Tech	USA	7	171	44	64	63
Univ of Sci & Tech China	China	8	157	20	87	50
Univ of Illinois at Urbana-Champaign	USA	9	153	70	27	56
Fudan Univ	China	10	140	49	65	26
Nanking Univ	China	11	138	42	68	28
Univ of Mumbai	India	12	136	55	23	58
Pennsylvania State Univ-Main Campus	USA	13	136	70	23	43
Univ of Michigan-Ann Arbor	USA	14	134	52	34	48
Shanghai Jiaotong Univ	China	15	133	8	27	98
Univ of Florida	USA	16	132	71	23	38
Nankai Univ	China	17	128	43	65	20
Univ of Wisconsin-Madison	USA	18	125	74	27	24
The Univ of Texas at Austin	USA	19	122	58	30	34
Univ of California-Davis	USA	20	119	75	29	15
Harvard Univ	USA	21	118	59	48	11
Brigham Young Univ	USA	22	116	52	39	25
Univ of California-Los Angeles	USA	23	116	61	38	17
Zhejiang Univ	China	24	115	9	31	75
China Univ of Sci and Tech	China	25	115	20	68	27
Total (incl. other institutions)			20,057	7,909	6,151	5,997

Source: Data from NORC(2008) based upon the NSF Survey of Earned Doctorates

Table 11: **Statistics of interest for the top 25 chemistry departments in our sample**

School	R & D exp.	total # students	# NSF fellows	Chinese (total)	Chinese (%)	
1	California Institute of Technology	36.9	186	26	16	8.6%
2	Harvard University	30.3	245	71	23	9.4%
3	University of California-Berkeley	29.5	449	37	17	3.8%
4	University of Illinois-Urbana	29.2	413	7	40	9.7%
5	Texas A&M University	24.7	315	2	45	14.3%
6	Georgia Institute of Technology	24.4	183	1	24	13.1%
7	University of California-San Diego	23.6	206	3	11	5.3%
8	University of California-Los Angeles	23.3	285	3	28	9.8%
9	University of Texas-Austin	22.7	273	4	31	11.4%
10	Rutgers	22.0	131	0	50	38.2%
11	Massachusetts Institute of Technology	21.5	258	41	20	7.8%
12	Northwestern University	21.3	228	5	34	14.9%
13	University of North Carolina-Chapel Hill	21.0	311	6	18	5.8%
14	Purdue University	20.4	288	1	50	17.4%
15	Pennsylvania State University	19.6	229	3	24	10.5%
16	Cornell University	19.2	210	4	22	10.5%
17	University of Washington-Seattle	18.6	192	2	11	5.7%
18	University of Colorado	18.4	199	2	6	3.0%
19	University of California-Irvine	17.9	216	0	15	6.9%
20	University of Wisconsin-Madison	17.9	445	15	30	6.7%
21	Stanford University	17.8	232	31	28	12.1%
22	Johns Hopkins University	17.1	99	0	10	10.1%
23	University of Michigan	16.6	264	2	34	12.9%
24	Louisiana State University	16.5	95	0	9	9.5%
25	Emory University	16.4	123	0	36	29.3%

Notes: R & D expenditures refers to 2007 R&D expenditures in chemistry in million USD. The number of students, Chinese students and NSF fellows are based on students graduating from chemistry departments between 1999 and 2008

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